

# STAT 464/864

## Midterm Practice Problems – Solutions, 2022

1. Suppose  $\{s_t\}$  is a seasonal component with period  $d = 2$ . Show that  $|\nabla^2 s_t|$  has no seasonal component, where  $\nabla$  is the difference operator.

**Solution:**  $\nabla^2 s_t = s_t - 2s_{t-1} + s_{t-2} = s_t - s_{t-1} - (s_{t-1} - s_{t-2})$ . But if  $\{s_t\}$  has period 2 then  $s_t - s_{t-1} = -(s_{t-1} - s_{t-2})$  and so  $\nabla^2 s_t = 2(s_t - s_{t-1})$ , and  $|\nabla^2 s_t| = 2|s_t - s_{t-1}|$ . This is the same for every  $t$  if  $\{s_t\}$  has period 2, and so has no seasonal component.

2. Let  $\{s_t\}$  and  $\{r_t\}$  be seasonal components with periods  $a$  and  $b$ , respectively, where  $a$  and  $b$  are distinct positive integers that do not share any prime factors. Let  $X_t = s_t + r_t + Z_t$ , where  $\{Z_t\}$  is a zero mean WN( $\sigma^2$ ) process. Find  $d_1$  and  $d_2$  such that  $\nabla_{d_1} \nabla_{d_2} X_t$  is stationary, where  $\nabla_d$  is the lag  $d$  difference operator, and find the ACF of  $\nabla_{d_1} \nabla_{d_2} X_t$ .

**Solution:** Take  $d_1 = a$  and  $d_2 = b$ . Then  $\nabla_a \nabla_b s_t = s_t - s_{t-b} - s_{t-a} + s_{t-(a+b)}$ . But  $s_t - s_{t-a} = 0$  and  $s_{t-b} - s_{t-b-a} = 0$ . Similarly,  $\nabla_a \nabla_b r_t = r_t - r_{t-b} - r_{t-a} + r_{t-(a+b)}$ . But  $r_t - r_{t-b} = 0$  and  $r_{t-a} - r_{t-a-b} = 0$ . So  $\nabla_a \nabla_b (s_t + r_t) = 0$  and  $\nabla_a \nabla_b X_t = \nabla_a \nabla_b Z_t = Z_t - Z_{t-a} - Z_{t-b} + Z_{t-(a+b)}$ . Then  $\nabla_a \nabla_b Z_t$  is the output of a linear filter  $\{\psi_j\}$  with  $\psi_0 = 1$ ,  $\psi_a = -1$ ,  $\psi_b = -1$ ,  $\psi_{a+b} = 1$ , and  $\psi_j = 0$  for  $j \notin \{0, a, b, a+b\}$ .  $\{\nabla_a \nabla_b X_t\}$  is thus a linear process, and hence is stationary. It is also causal and has ACVF  $\gamma_X(h) = \sum_{j=0}^{\infty} \psi_j \psi_{j+h}$ . For  $h \geq 0$ , the only  $h$  for which  $\gamma_X(h)$  is nonzero are  $h = 0, a, b, |b-a|$ , and  $a+b$ . For these  $h$  we have  $\gamma_X(0) = 4\sigma^2$ ,  $\gamma_X(a) = -2\sigma^2$ ,  $\gamma_X(b) = -2\sigma^2$ ,  $\gamma_X(|b-a|) = \sigma^2$ , and  $\gamma_X(a+b) = \sigma^2$ . Finally,  $\gamma_X(-h) = \gamma_X(h)$ . The ACF of  $\{X_t\}$  is then given by  $\rho_X(0) = 1$ ,  $\rho_X(\pm a) = \rho_X(\pm b) = -\frac{1}{2}$ ,  $\rho_X(\pm|b-a|) = \frac{1}{4}$ ,  $\rho_X(\pm(a+b)) = \frac{1}{4}$ , and  $\rho_X(h) = 0$  otherwise.

3. Suppose  $\{X_t\}$  and  $\{Y_t\}$  are independent (i.e.,  $X_r$  and  $Y_s$  are independent for every  $r$  and  $s$ ), 0-mean, stationary processes with autocovariance functions  $\gamma_X(h)$  and  $\gamma_Y(h)$ , respectively. Let  $Z_t = X_t Y_t$ . What is the ACVF and ACF of  $\{Z_t\}$ ?

**Solution:** Since  $E[Z_t] = E[X_t Y_t] = E[X_t]E[Y_t] = 0$  for all  $t$  the ACVF of  $\{Z_t\}$  at lag  $h$  is  $\gamma_Z(h) = E[Z_t Z_{t+h}] = E[X_t Y_t X_{t+h} Y_{t+h}] = E[X_t X_{t+h}]E[Y_t Y_{t+h}] = \gamma_X(h)\gamma_Y(h)$ . Then the ACF of  $\{Z_t\}$  is  $\rho_Z(h) = \rho_X(h)\rho_Y(h)$ , where  $\rho_X(h)$  and  $\rho_Y(h)$  are the ACFs of  $\{X_t\}$  and  $\{Y_t\}$ , respectively, at lag  $h$ .

4. For  $d$  odd, give a symmetric linear filter (i.e.,  $a_{-j} = a_j$ ) that eliminates seasonal components with period  $d$  and passes linear trends. What if  $d$  is even?

**Solution:** For  $d$  odd set  $a_j = \frac{1}{d}$  for  $j = -\frac{d-1}{2}, \dots, \frac{d-1}{2}$ , and  $a_j = 0$  otherwise. For  $d$  even set  $a_j = \frac{1}{2d}$  for  $j = -d, \dots, -1$  and  $j = 1, \dots, d$ , and set  $a_j = 0$  otherwise.

5. Consider the exponential smoother with smoothing parameter  $\alpha \in (0, 1)$  and let  $\{X_t\}$  be the result of applying this smoother to  $\{Z_t\}$ , where  $\{Z_t\}$  is a zero-mean WN( $\sigma^2$ ) process, i.e.,  $X_t = \sum_{i=0}^{\infty} \alpha(1-\alpha)^i Z_{t-i}$ . Find the autocovariance function (ACVF) of  $\{X_t\}$  at all lags  $h$ .

**Solution:**  $\{X_t\}$  is the output of the linear filter  $\{\psi_j\}$ , where,  $\psi_j = \alpha(1-\alpha)^j$  for  $j \geq 0$  and  $\psi_j = 0$  for  $j < 0$ . For  $h \geq 0$ , the ACVF of  $\{X_t\}$  is

$$\gamma_X(h) = \sum_{j=0}^{\infty} \alpha(1-\alpha)^j \alpha(1-\alpha)^{j+h} = \alpha^2(1-\alpha)^h \frac{1}{1-(1-\alpha)^2} = (1-\alpha)^h \frac{\alpha}{2-\alpha},$$

and  $\gamma_X(-h) = \gamma_X(h)$ .

6. Let  $X_t = Y_0 s_t + Y_1 s_{t-1}$ , where  $\{s_t\}$  is seasonal with period 2, and  $Y_t = Z_t + \theta Z_{t-1}$  is an MA(1) process with MA coefficient  $\theta$ , and  $\{Z_t\}$  is a zero-mean WN( $\sigma^2$ ) process. Compute the ACF of  $\{X_t\}$ .

**Solution:**  $\{X_t\}$  is clearly zero-mean. So for  $\text{Cov}(X_t, X_{t+h}) = E[X_t, X_{t+h}] = E[(Y_0 s_t + Y_1 s_{t-1})(Y_0 s_{t+h} + Y_1 s_{t+h-1})] = s_t s_{t+h} E[Y_0^2] + s_{t-1} s_{t+h-1} E[Y_1^2] + (s_t s_{t+h-1} + s_{t-1} s_{t+h}) E[Y_0 Y_1] = \gamma_Y(0)(s_t s_{t+h} + s_{t-1} s_{t+h-1}) + \gamma_Y(1)(s_t s_{t+h-1} + s_{t-1} s_{t+h})$ , where  $\gamma_Y(h)$  is the ACVF of  $\{Y_t\}$ . If  $h$  is even then  $s_t s_{t+h} + s_{t-1} s_{t+h-1} = s_0^2 + s_1^2$  and  $s_t s_{t+h-1} + s_{t-1} s_{t+h} = 2s_0 s_1$  for every  $t$ , while if  $h$  is odd then  $s_t s_{t+h} + s_{t-1} s_{t+h-1} = 2s_0 s_1$  and  $s_t s_{t+h-1} + s_{t-1} s_{t+h} = s_0^2 + s_1^2$  for every  $t$ . Therefore,  $\text{Cov}(X_t, X_{t+h})$  does not depend on  $t$  and so  $\{X_t\}$  is stationary. For the MA(1) process with parameter  $\theta$ ,  $\gamma_Y(0) = \sigma^2(1+\theta^2)$  and  $\gamma_Y(1) = \sigma^2\theta$ . For  $h$  even the ACVF of  $\{X_t\}$  is then  $\gamma_X(h) = \sigma^2(1+\theta^2)(s_0^2 + s_1^2) + 2\sigma^2\theta s_0 s_1$  and for  $h$  odd  $\gamma_X(h) = 2\sigma^2(1+\theta^2)s_0 s_1 + \sigma^2\theta(s_0^2 + s_1^2)$ . Then the ACF of  $\{X_t\}$  is  $\rho_X(h) = \frac{\gamma_X(h)}{\gamma_X(0)} = 1$  if  $h$  is even and equals  $\frac{2(1+\theta^2)s_0 s_1 + \theta(s_0^2 + s_1^2)}{(1+\theta^2)(s_0^2 + s_1^2) + 2\theta s_0 s_1}$  if  $h$  is odd.

7. Show that the two MA(1) processes

$$\begin{aligned} X_t &= Z_t + \theta Z_{t-1}, & \{Z_t\} &\sim \text{WN}(0, \sigma^2) \\ Y_t &= \tilde{Z}_t + \frac{1}{\theta} \tilde{Z}_{t-1}, & \{\tilde{Z}_t\} &\sim \text{WN}(0, \sigma^2 \theta^2), \end{aligned}$$

where  $0 < |\theta| < 1$ , have the same autocovariance function.

**Solution:** The ACVF of  $\{X_t\}$  is  $\gamma_X(0) = \sigma^2(1+\theta^2)$ ,  $\gamma_X(\pm 1) = \sigma^2\theta$ , and  $\gamma_X(h) = 0$  otherwise. The ACVF of  $\{Y_t\}$  is  $\gamma_Y(0) = \sigma^2\theta^2(1+1/\theta^2) = \sigma^2(\theta^2 + 1)$ ,  $\gamma_Y(\pm 1) = \sigma^2\theta^2(1/\theta) = \sigma^2\theta$ , and  $\gamma_Y(h) = 0$  otherwise. Therefore,  $\{X_t\}$  and  $\{Y_t\}$  have the same ACVF.